

Comparing the Estimates of Mixed Models for Binary or Ordinal Data

October 20, 2006
Meeting of the Society of Multivariate
Experimental Psychology

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Mixed Models

- Mixed models provide a model-based way to account for dependence in clustered or longitudinal data.
- Dependence is usually accommodated through the introduction of random effects.
 - Random intercepts to account for mean differences between units.

The Linear Mixed Model

- The general form of the linear mixed model is

$$y_{ij} = \mathbf{x}'_{ij}\boldsymbol{\beta} + \mathbf{z}'_{ij}\mathbf{u}_j + r_{ij}$$

- observation i is nested within unit j .
 - \mathbf{x}_{ij} is a vector of predictors with fixed effects.
 - $\boldsymbol{\beta}$ is the vector of fixed effects.
 - \mathbf{z}_{ij} is a vector of predictors with random effects.
 - \mathbf{u}_j is the vector of random effects for unit j .
 - r_{ij} is the observation-specific residual.
- Typically assume that

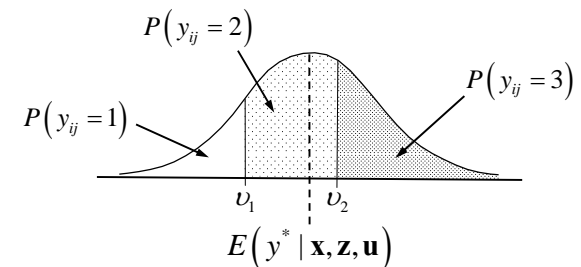
$$\begin{pmatrix} \mathbf{u} \\ \mathbf{r} \end{pmatrix} \sim N \left(\begin{bmatrix} \mathbf{0} \\ \mathbf{0} \end{bmatrix}, \begin{bmatrix} \boldsymbol{\Sigma}_u & 0 \\ 0 & \sigma_r^2 \mathbf{I}_n \end{bmatrix} \right)$$

Mixed Model for Binary and Ordinal Outcomes

- Assume that y represents a coarse version of a continuous underlying variable y^*

$$y_{ij}^* = \mathbf{x}'_{ij}\boldsymbol{\beta} + \mathbf{z}'_{ij}\mathbf{u}_j + r_{ij}$$

$$y_{ij} = c \text{ if } v^{c-1} < y_{ij}^* \leq v^c \quad (v_0 \equiv -\infty; v_C \equiv \infty)$$



Scaling for Binary and Ordinal Outcomes

- Because y^* is not directly observed, its location and scale are arbitrary.
- Set the location by fixing first threshold: $\nu_1 \equiv 0$
- Set the scale by fixing σ_r^2
 - In the probit model, assume $r_{ij} \sim N(0,1)$
 - In the logit model, assume $r_{ij} \sim \text{logistic}\left(0, \frac{\pi^2}{3}\right)$

Comparing Sequential Models

- Can we compare the estimates of sequential models?
 - Model 1: $y_{ij}^* = \beta_0 + u_j + r_{ij}$
 - Model 2: $y_{ij}^* = \beta_0 + \beta_1 x_{1ij} + u_j + r_{ij}$
 - Model 3: $y_{ij}^* = \beta_0 + \beta_1 x_{1ij} + \beta_2 x_{2ij} + u_j + r_{ij}$
- Across these models, σ_r^2 remains constant.
 - In the probit model: $\sigma_r^2 = 1$
 - In the logit model: $\sigma_r^2 = \pi^2 / 3$
- Variance of y^* must then change between models, putting estimates on different scales.

Placing Estimates on a Commensurate Scale

- To compare the estimates of simple logistic or probit models, Winship & Mare (1983, 1984) recommended equating the model-implied marginal variance of y^* across the fitted models.

Placing Estimates on a Commensurate Scale

- The model-implied marginal variance of y^* in the mixed model is

$$V(y_{ij}^*) = \beta' \Sigma_x \beta + \mu_z' \Sigma_u \mu_z + \text{VEC}(\Sigma_z)' \text{VEC}(\Sigma_u) + \sigma_r^2$$

- It then follows that

$$\left(\beta' \Sigma_x \beta + \mu_z' \Sigma_u \mu_z + \text{VEC}(\Sigma_z)' \text{VEC}(\Sigma_u) + \sigma_r^2 \right)^{-1} V(y_{ij}^*) = 1$$

- To rescale the implied variance to equal a , note that

$$a \left(\beta' \Sigma_x \beta + \mu_z' \Sigma_u \mu_z + \text{VEC}(\Sigma_z)' \text{VEC}(\Sigma_u) + \sigma_r^2 \right)^{-1} V(y_{ij}^*) = a$$

Placing Estimates on a Commensurate Scale

- Recall that $s^2 V(y_{ij}^*) = V(sy_{ij}^*)$, where s is a scaling factor

$$a \underbrace{\left(\boldsymbol{\beta}' \boldsymbol{\Sigma}_x \boldsymbol{\beta} + \boldsymbol{\mu}'_z \boldsymbol{\Sigma}_u \boldsymbol{\mu}_z + \text{VEC}(\boldsymbol{\Sigma}_z)' \text{VEC}(\boldsymbol{\Sigma}_u) + \sigma_r^2 \right)^{-1}}_{s^2} V(y_{ij}^*) = a$$

- We can apply this to the mixed model equation as follows:

$$s = a^{1/2} \left(\boldsymbol{\beta}' \boldsymbol{\Sigma}_x \boldsymbol{\beta} + \boldsymbol{\mu}'_z \boldsymbol{\Sigma}_u \boldsymbol{\mu}_z + \text{VEC}(\boldsymbol{\Sigma}_z)' \text{VEC}(\boldsymbol{\Sigma}_u) + \sigma_r^2 \right)^{-1/2}$$

$$sy_{ij}^* = \mathbf{x}'_{ij} (s\boldsymbol{\beta}) + \mathbf{z}'_{ij} (s\mathbf{u}_j) + (sr_{ij})$$

$$\text{VAR}(sy_{ij}^*) = a$$

Comparing Fitted Models

- Given the equation

$$sy_{ij}^* = \mathbf{x}'_{ij} (s\boldsymbol{\beta}) + \mathbf{z}'_{ij} (s\mathbf{u}_j) + (sr_{ij})$$

the rescaled fixed effects are $s\boldsymbol{\beta}$

the rescaled covariance parameters are $s^2 \boldsymbol{\Sigma}_u$ and $s^2 \sigma_r^2$

- For any two models to be compared, can rescale estimates using a common a value so that the units are commensurate.

Simulated Data Example

- 10,000 Clusters, 1 to 10 observations per cluster
- Population Model

$$y_{ij}^* = \beta_0 + \beta_1 x_{1ij} + \beta_2 x_{2ij} + u_j + r_{ij} \quad y = 1 \text{ if } y_{ij}^* > 0; \text{ else } y = 0$$

$$\beta_0 = 0; \beta_1 = \beta_2 = 1$$

$$\sigma_u^2 = .05; \sigma_r^2 = .15; r_{ij} \sim N$$

$$\boldsymbol{\mu}_x = \mathbf{0}$$

$$\boldsymbol{\Sigma}_x = \begin{pmatrix} .1+.3 & 0 \\ 0 & .1+.3 \end{pmatrix}$$

$$\text{CORR}(x_1, x_2) = 0$$

Sequence of Models

- Fit two mixed-effects probit models to y
 - Model 1: $P(y_{ij} = 1) = g(\beta_0 + \beta_1 x_{1ij} + u_j)$
 - Model 2: $P(y_{ij} = 1) = g(\beta_0 + \beta_1 x_{1ij} + \beta_2 x_{2ij} + u_j)$

g^{-1} chosen to be probit link function.
- For comparison, fit two linear mixed models to simulated y^*
 - Model 1: $y_{ij}^* = \beta_0 + \beta_1 x_{1ij} + u_j + r_{ij}$
 - Model 2: $y_{ij}^* = \beta_0 + \beta_1 x_{1ij} + \beta_2 x_{2ij} + u_j + r_{ij}$

Estimates for models fit to simulated data.

Effects	Fit to y^*		Fit to y	
	Linear	Probit	Rescaled ($a=1$)	Rescaled ($a=2$)
<i>Model 1</i>				
β_0	0.01	0.01	0.01	0.01
β_1	1.02	1.52	1.03	1.45
σ_u^2	0.15	0.29	0.13	0.26
σ_r^2	0.44	1.00	0.46	0.91
<i>Model 2</i>				
β_0	0.00	0.01	0.00	0.01
β_1	1.00	2.61	1.00	1.42
β_2	0.99	2.62	1.00	1.42
σ_u^2	0.05	0.30	0.04	0.09
σ_r^2	0.15	1.00	0.15	0.29

Conclusions

- For continuous outcomes, often recommended to fit and compare sequentially fit models.
- For binary and ordinal outcomes, sequentially fit models cannot be meaningfully compared unless the estimates are rescaled to equate the model-implied variance of y^* .

Of Additional Interest...

- Attention to scaling also clarifies several other issues that arise with mixed models for binary/ordinal outcomes:
 - Why marginal and conditional model estimates differ (inclusion of random effects changes scale)
 - Why estimators (e.g., MQL, PQL) that produce biased variance component estimates also produce biased fixed effects estimates (on wrong scale)
 - Why misspecification of the variance component structure can bias the fixed effects estimates (on wrong scale)